

Sol: (CNC 2014) (BCPST)

On a $X \sim \mathcal{E}(\lambda)$; f sa densité. ($\lambda > 0$)

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & (\text{si } x \geq 0) \\ 0 & (\text{si } x < 0) \end{cases}$$

$$Y = [X]$$

Ainsi: Y a valeurs dans \mathbb{Z}

$$\forall k \in \mathbb{Z}, (Y = k) = (k \leq X < k+1)$$

1) 1) Mq Y a valeurs dans \mathbb{N} .

\perp Il suffit de mpr $X \geq 0$.

Caol - $P(X < 0) = 0$

Dabei:

$$P(X < 0) = P(X \leq 0) = \int_{-\infty}^0 f(x) dx = 0$$

car $(\forall x < 0, f(x) = 0)$

1) 2) Sei $k \in \mathbb{N}^*$, $P(Y = k-1) = ?$

$$\begin{aligned} P(Y = k-1) &= P(\lfloor X \rfloor = k-1) \\ &= P(k-1 \leq X < k) \quad \parallel \int_{k-1}^k f(x) dx \\ &= \int_{k-1}^k f(x) dx = \int_{k-1}^k \lambda e^{-\lambda x} dx \end{aligned}$$

$$P(a \leq X \leq b) = F(b) - F(a) = \int_a^b f(x) dx$$

Rappel

$$\forall k \in \mathbb{N}^*, P(Y=k-1) = e^{-\lambda k} (e^\lambda - 1)$$

1)3) a) $(Y+1)$ suit une loi géométrique ; en effet :

(i) $(Y+1)(\Omega) = \mathbb{N}^*$ (car $Y(\Omega) = \mathbb{N}$)

(ii) Soit $k \in \mathbb{N}^*$,

$$P(Y+1=k) = P(Y=k-1)$$

$$= e^{-\lambda k} (e^\lambda - 1)$$

$$= (e^{-\lambda})^k (e^\lambda - 1)$$

$$= (e^{-\lambda})^{k-1} \cdot e^{-\lambda} (e^\lambda - 1)$$

$$= (e^{-\lambda})^{k-1} (1 - e^{-\lambda})$$

$$\forall c: (Y+1) \sim G(1 - e^{-\lambda})$$

1) 3) b) (i) On a $(Y+1) \sim G(p)$, où $p = 1 - e^{-\lambda}$

$$E(Y+1) = \frac{1}{p} = \frac{1}{1 - e^{-\lambda}}$$

$$V(Y+1) = \frac{q}{p^2} = \frac{e^{-\lambda}}{(1 - e^{-\lambda})^2}$$

Rappel. Si $X \sim G(p)$, on a $E(X) = \frac{1}{p}$ et $V(X) = \frac{q}{p^2}$
où $q = 1 - p$.

$$(ii) E(Y) = E(Y+1) - 1 = \frac{e^{-\lambda}}{1 - e^{-\lambda}}$$

$$V(Y) = V(Y+1) = \frac{e^{-\lambda}}{(1 - e^{-\lambda})^2} \quad V(aY+b) = a^2 V(Y)$$

$$2) Z = X - Y; \text{ où } Y = [X]$$

$$2)1) Z(\Omega) = [0, 1[$$

$$2)2) \text{ Soit } x \in [0, 1[. \text{ On a;}$$

$$P(Z \leq x) = \sum_{k=0}^{+\infty} P(Z \leq x, Y=k) \quad \left(\begin{array}{l} \text{car } (Y=k) \\ \text{est un système} \\ \text{complet d'événements} \end{array} \right)$$

Fermele des probas totales

$$= \sum_{k=0}^{+\infty} P(X - Y \leq x, Y=k) \quad \begin{array}{l} \text{et} \\ \uparrow \end{array}$$

$$= \sum_{k=0}^{+\infty} P(X - k \leq x, Y=k)$$

$$= \sum_{k=0}^{+\infty} P(X \leq x+k, k \leq X < k+1) \quad \begin{array}{l} \text{et} \\ \uparrow \end{array} \quad \left| \begin{array}{l} [X] = k \\ \parallel \\ k \leq X < k+1 \end{array} \right.$$

Or $(x \leq n+k, \overset{+e+}{k} \leq x < k+1) = (k \leq x \leq n+k)$

! Ca $x+k < k+1$

! $P(Z \leq x) = \sum_{k=0}^{+\infty} P(k \leq X \leq n+k)$

$= \sum_{k=0}^{+\infty} \left(\int_k^{n+k} f(t) dt \right)$

$= \sum_{k=0}^{+\infty} \left(\int_k^{n+k} \lambda e^{-\lambda t} dt \right)$

$= \sum_{k=0}^{+\infty} \left[-e^{-\lambda t} \right]_k^{n+k}$

$P(a \leq X \leq b)$
 \parallel
 $\int_a^b f(t) dt$

$$= \sum_{k=0}^{+\infty} (e^{-\lambda k} - e^{-\lambda(\lambda+k)})$$

$$= \sum_{k=0}^{+\infty} e^{-\lambda k} (1 - e^{-\lambda})$$

$$= (1 - e^{-\lambda}) \cdot \sum_{k=0}^{+\infty} (e^{-\lambda})^k$$

$$= \frac{1 - e^{-\lambda x}}{1 - e^{-\lambda}}$$

GF(D)

$$0 < e^{-\lambda} < 1$$

(as $\lambda > 0$)

$$\sum_{k=0}^{+\infty} q^k = \frac{1}{1-q}$$

$$\text{s.t. } |q| < 1$$

2)3) $E(Z) = ?$

$$E(Z) = \int_{-\infty}^{+\infty} t f_Z(t) dt$$

Or:

$$\left(\forall x \in [0, 1[, F_Z(x) = \frac{1 - e^{-\lambda x}}{1 - e^{-\lambda}} \right)$$

$$Z(\Omega) = [0, 1[$$

$$\forall x < 0, F_Z(x) = P(Z \leq x) = P(\emptyset) = 0$$

$$\forall x \geq 1, F_Z(x) = P(Z \leq x) = P(Z < 1) = 1$$

$f_Z(x) =$	0	$x < 0$
	$\frac{\lambda e^{-\lambda x}}{1 - e^{-\lambda}}$	$0 < x < 1$
	0	$x > 1$

Ans:

$$E(z) = \int_0^1 t f_2(t) dt$$

$$= \frac{1}{1-e^{-\lambda}} \int_0^1 \lambda + e^{-\lambda t} dt$$

$$= \frac{1}{1-e^{-\lambda}} \int_0^1 (-e^{-\lambda t})' t dt$$

$$= \frac{1}{1-e^{-\lambda}} \left(\underbrace{\left[-e^{-\lambda t} t \right]_0^1}_{=-e^{-\lambda}} + \underbrace{\int_0^1 e^{-\lambda t} dt}_{=\frac{1-e^{-\lambda}}{\lambda}} \right)$$

$$= \frac{1}{1-e^{-\lambda}} \left(\frac{1-e^{-\lambda}}{\lambda} - e^{-\lambda} \right)$$

$$E(Z) = \frac{1}{\lambda} - \frac{e^{-\lambda}}{1-e^{-\lambda}}$$

Fin

RIR: On peut calculer $E(Z)$ par linéarité de l'espérance:

$$\begin{aligned} E(Z) &= E(X-Y) = E(X) - E(Y) \\ &= \frac{1}{\lambda} - \frac{e^{-\lambda}}{1-e^{-\lambda}} \end{aligned}$$

